

The FiniteSet Independence Criterion (FSIC)

Wittawat Jitkrittum

Zoltán Szabó

Arthur Gretton

Gatsby Unit

University College London

wittawat@gatsby.ucl.ac.uk

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What Is Independence Testing?

- Let $(X, Y) \in \mathbb{R}^{d_x} \times \mathbb{R}^{d_y}$ be random vectors following P_{xy} .
- Given a joint sample $\{(\mathbf{x}_i, \mathbf{y}_i)\}_{i=1}^n \sim P_{xy}$ (unknown), test

$$H_0 : P_{xy} = P_x P_y,$$

$$\text{vs. } H_1 : P_{xy} \neq P_x P_y.$$

- Compute a test statistic $\hat{\lambda}_n$. Reject H_0 if $\hat{\lambda}_n > T_\alpha$ (threshold).
- $T_\alpha = (1 - \alpha)$ -quantile of the null distribution.

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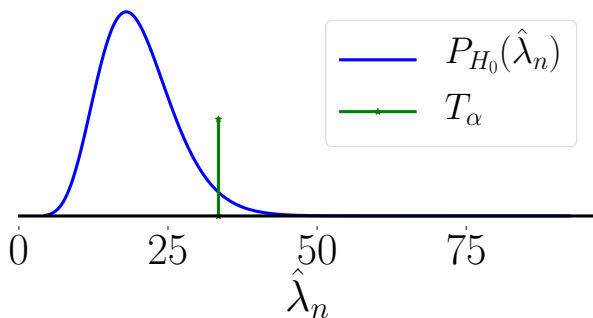
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Motivations

Modern state-of-the-art test is HSIC [Gretton et al., 2005].

- ✓ **Nonparametric** i.e., no assumption on P_{xy} . Kernel-based.
- ✗ **Slow**. Runtime: $\mathcal{O}(n^2)$ where n = sample size.
- ✗ **No systematic way to choose kernels**.

Propose the **Finite-Set Independence Criterion (FSIC)**.

- 1 **Nonparametric**.
- 2 **Linear-time**. Runtime complexity: $\mathcal{O}(n)$. Fast.
- 3 **Tunable** i.e., well-defined criterion for parameter tuning.

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Proposal: The Finite-Set Independence Criterion (FSIC)

1 Pick 2 positive definite kernels: k for X , and l for Y .

- Gaussian kernel: $k(\mathbf{x}, \mathbf{v}) = \exp\left(-\frac{\|\mathbf{x}-\mathbf{v}\|^2}{2\sigma_x^2}\right)$.

2 Pick some **feature** $(\mathbf{v}, \mathbf{w}) \in \mathbb{R}^{d_x} \times \mathbb{R}^{d_y}$

3. Transform $(\mathbf{x}, \mathbf{y}) \mapsto (k(\mathbf{x}, \mathbf{v}), l(\mathbf{y}, \mathbf{w}))$ then measure covariance

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$$\text{FSIC}^2(X, Y) = \text{cov}_{(\mathbf{x}, \mathbf{y}) \sim P_{xy}}^2 [k(\mathbf{x}, \mathbf{v}), l(\mathbf{y}, \mathbf{w})] .$$

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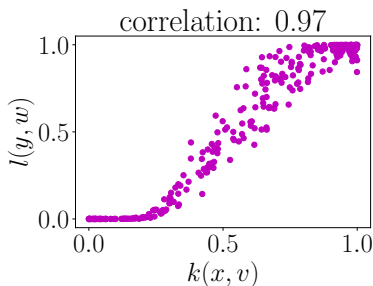
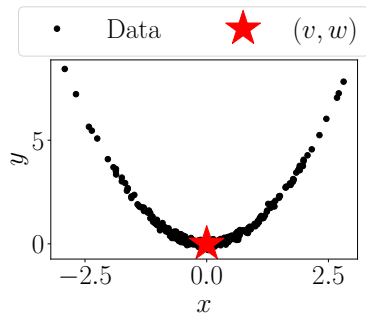
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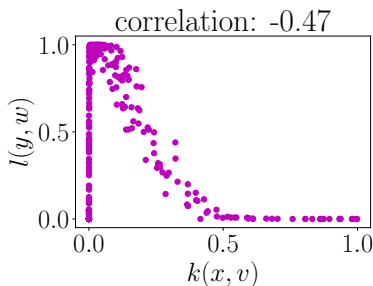
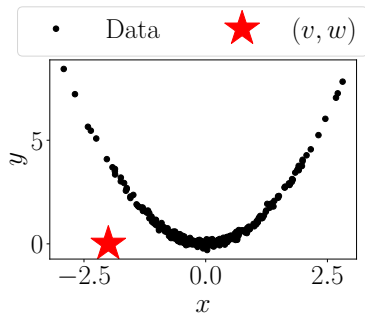
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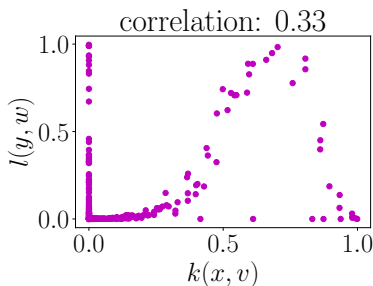
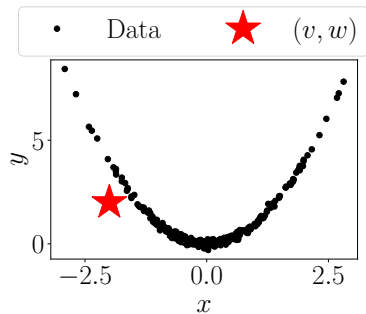
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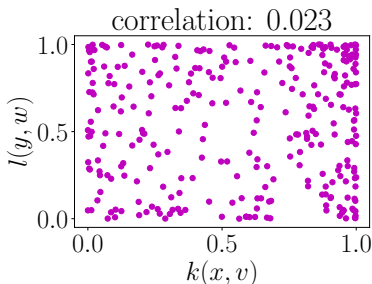
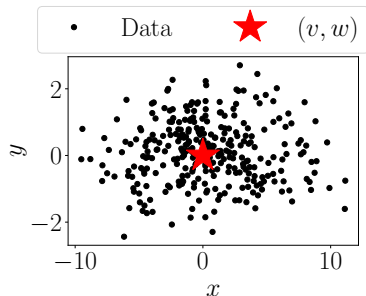
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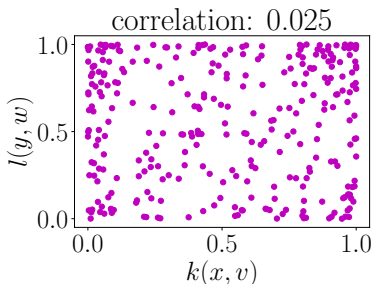
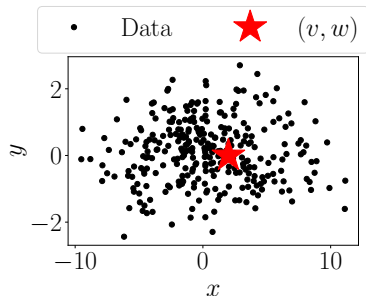
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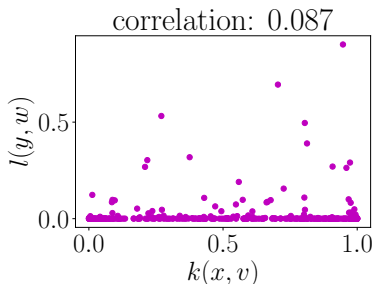
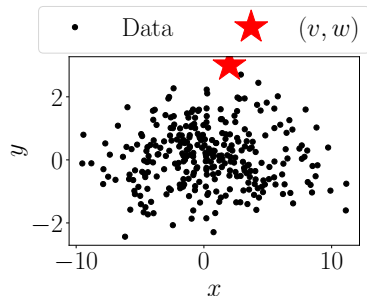
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General Form of FSIC

$$\text{FSIC}^2(X, Y) = \frac{1}{J} \sum_{j=1}^J \text{cov}_{(\mathbf{x}, \mathbf{y}) \sim P_{xy}}^2 [k(\mathbf{x}, \mathbf{v}_j), l(\mathbf{y}, \mathbf{w}_j)],$$

for J **features** $\{(\mathbf{v}_j, \mathbf{w}_j)\}_{j=1}^J \in \mathbb{R}^{d_x} \times \mathbb{R}^{d_y}$.

Proposition 1.

Assume

- 1 Kernels k and l satisfy some conditions (e.g. Gaussian kernels).
- 2 Features $\{(\mathbf{v}_i, \mathbf{w}_i)\}_{i=1}^J$ are drawn from a distribution with a density.

Then, for any $J \geq 1$,

$\text{FSIC}(X, Y) = 0$ if and only if X and Y are independent

Under $H_0 : P_{xy} = P_x P_y$,

$n \widehat{\text{FSIC}}^2 \sim$ weighted sum of J **dependent** χ^2 variables.

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Normalized FSIC (NFSIC)

- Let $\hat{\mathbf{u}} := \left(\widehat{\text{cov}}[k(\mathbf{x}, \mathbf{v}_1), l(\mathbf{y}, \mathbf{w}_1)], \dots, \widehat{\text{cov}}[k(\mathbf{x}, \mathbf{v}_J), l(\mathbf{y}, \mathbf{w}_J)] \right)^\top \in \mathbb{R}^J$.
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$$\widehat{\text{NFSIC}}^2(X, Y) = \hat{\lambda}_n := n \hat{\mathbf{u}}^\top \left(\hat{\Sigma} + \gamma_n \mathbf{I} \right)^{-1} \hat{\mathbf{u}},$$

with a regularization parameter $\gamma_n \geq 0$.

- $\hat{\Sigma}_{ij}$ = covariance of \hat{u}_i and \hat{u}_j .

Theorem 1 (NFSIC test is consistent).

Assume $\gamma_n \rightarrow 0$, and same conditions on k and l as before.

- 1 Under H_0 , $\hat{\lambda}_n \xrightarrow{d} \chi^2(J)$ as $n \rightarrow \infty$. Easy to get threshold T_α .
- 2 Under H_1 , $\mathbb{P}(\text{reject } H_0) \rightarrow 1$ as $n \rightarrow \infty$.

- Complexity: $\mathcal{O}(J^3 + J^2 n + (d_x + d_y) J n)$. Only need small J .

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Tuning Features and Kernels

- Split the data into training (**tr**) and test (**te**) sets.

Procedure:

- 1 Choose $\{(\mathbf{v}_i, \mathbf{w}_i)\}_{i=1}^J$ and Gaussian widths by maximizing $\hat{\lambda}_n^{(\text{tr})}$ (i.e., computed on the training set). Gradient ascent.
- 2 Reject H_0 if $\hat{\lambda}_n^{(\text{te})} > (1 - \alpha)$ -quantile of $\chi^2(J)$.

- Splitting avoids overfitting.

Theorem 2.

- *This procedure increases a lower bound on $\mathbb{P}(\text{reject } H_0 \mid H_1 \text{ true})$ (test power).*
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Simulation Settings

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	Method	Description
1	NFSIC-opt	NFSIC with optimization. $\mathcal{O}(n)$.
2	QHSIC [Gretton et al., 2005]	State-of-the-art HSIC. $\mathcal{O}(n^2)$.
3	NFSIC-med	NFSIC with random features.
4	NyHSIC	Linear-time HSIC with Nystrom approx.
5	FHSIC	Linear-time HSIC with random Fourier features
6	RDC [Lopez-Paz et al., 2013]	Canonical Correlation Analysis with cosine basis.

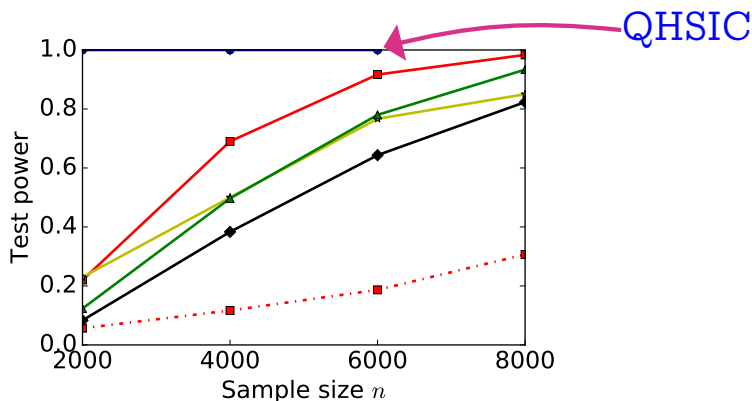


■—■ NFSIC-opt ■...■ NFSIC-med ●—● QHSIC ★—★ NyHSIC ◆—◆ FHSIC ▲—▲ RDC

- $J = 10$ in NFSIC.

Youtube Video (X) vs. Caption (Y).

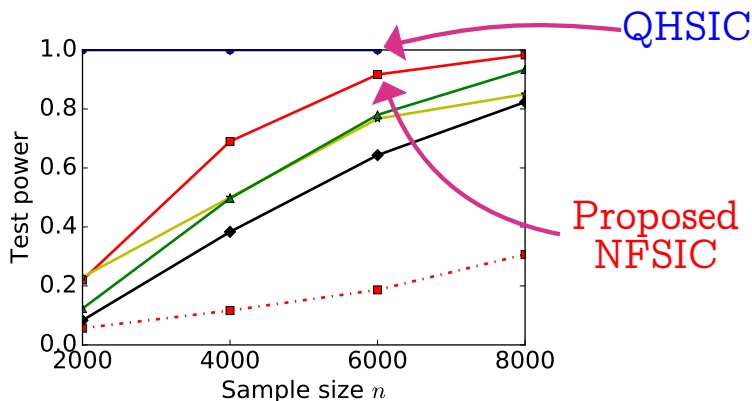
- $X \in \mathbb{R}^{2000}$: Fisher vector encoding of motion boundary histograms descriptors [Wang and Schmid, 2013].
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■ For large n , NFSIC is comparable to HSIC.

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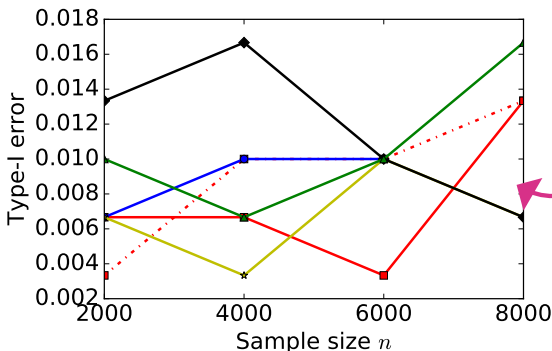
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Conclusions

- Proposed **The Finite Set Independence Criterion (FSIC)**.
 - Independence test based on FSIC is
 - 1 nonparametric,
 - 2 linear-time,
 - 3 adaptive (parameters automatically tuned).
-

An Adaptive Test of Independence with Analytic Kernel Embeddings
Wittawat Jitkrittum, Zoltán Szabó, Arthur Gretton
<https://arxiv.org/abs/1610.04782>
(to appear in ICML 2017)

- Python code: <https://github.com/wittawatj/fsic-test>

Questions?

Thank you

Reference

Coauthors:



Zoltán Szabó
École Polytechnique



Arthur Gretton
Gatsby Unit, UCL

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Requirements on the Kernels

Definition 1 (Analytic kernels).

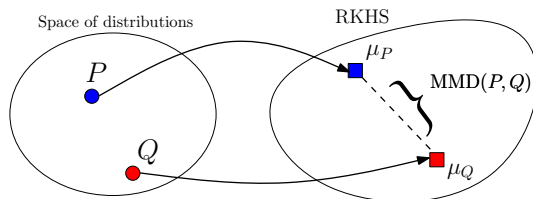
$k : \mathcal{X} \times \mathcal{X} \rightarrow \mathbb{R}$ is said to be analytic if for all $\mathbf{x} \in \mathcal{X}$, $\mathbf{v} \mapsto k(\mathbf{x}, \mathbf{v})$ is a real analytic function on \mathcal{X} .

- Analytic: Taylor series about \mathbf{x}_0 converges for all $\mathbf{x}_0 \in \mathcal{X}$.
- $\implies k$ is infinitely differentiable.

Definition 2 (Characteristic kernels).

- Let $\mu_P(\mathbf{v}) := \mathbb{E}_{\mathbf{z} \sim P}[k(\mathbf{z}, \mathbf{v})]$.

k is said to be characteristic if μ_P is unique for distinct P . Equivalently, $P \mapsto \mu_P$ is injective.



Optimization Objective = Power Lower Bound

- Recall $\hat{\lambda}_n := n\hat{\mathbf{u}}^\top (\hat{\Sigma} + \gamma_n \mathbf{I})^{-1} \hat{\mathbf{u}}$.
- Let $\text{NFSIC}^2(X, Y) := \lambda_n := n\mathbf{u}^\top \Sigma^{-1} \mathbf{u}$.

Theorem 3 (A lower bound on the test power).

- 1 With some conditions, the test power $\mathbb{P}_{H_1}(\hat{\lambda}_n \geq T_\alpha) \geq L(\lambda_n)$ where

$$L(\lambda_n) = 1 - 62e^{-\xi_1 \gamma_n^2 (\lambda_n - T_\alpha)^2 / n} - 2e^{-[0.5n](\lambda_n - T_\alpha)^2 / [\xi_2 n^2]} \\ - 2e^{-[(\lambda_n - T_\alpha) \gamma_n (n-1) / 3 - \xi_3 n - c_3 \gamma_n^2 n(n-1)]^2 / [\xi_4 n^2 (n-1)]},$$

where $\xi_1, \dots, \xi_4, c_3 > 0$ are constants.

- 2 For large n , $L(\lambda_n)$ is increasing in λ_n .

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Set test locations and Gaussian widths = $\arg \max L(\lambda_n) = \arg \max \lambda_n$

An Estimator of $\widehat{\text{NFSIC}}^2$

$$\hat{\lambda}_n := n \hat{\mathbf{u}}^\top \left(\hat{\Sigma} + \gamma_n \mathbf{I} \right)^{-1} \hat{\mathbf{u}},$$

- J test locations $\{(\mathbf{v}_i, \mathbf{w}_i)\}_{i=1}^J \sim \eta$.
- $\mathbf{K} = [k(\mathbf{v}_i, \mathbf{x}_j)] \in \mathbb{R}^{J \times n}$
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Estimators

- 1 $\hat{\mathbf{u}} = \frac{(\mathbf{K} \circ \mathbf{L}) \mathbf{1}_n}{n-1} - \frac{(\mathbf{K} \mathbf{1}_n) \circ (\mathbf{L} \mathbf{1}_n)}{n(n-1)}.$
 - 2 $\hat{\Sigma} = \frac{\Gamma \Gamma^\top}{n}$ where $\Gamma := (\mathbf{K} - n^{-1} \mathbf{K} \mathbf{1}_n \mathbf{1}_n^\top) \circ (\mathbf{L} - n^{-1} \mathbf{L} \mathbf{1}_n \mathbf{1}_n^\top) - \hat{\mathbf{u}} \mathbf{1}_n^\top$.
- $\hat{\lambda}_n$ can be computed in $\mathcal{O}(J^3 + J^2 n + (d_x + d_y) J n)$ time.

Main Point: Linear in n . Cubic in J (small).

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Alternative View of the Witness $u(\mathbf{v}, \mathbf{w})$

The witness $u(\mathbf{v}, \mathbf{w})$ can be rewritten as

$$\begin{aligned}u(\mathbf{v}, \mathbf{w}) &:= \mu_{xy}(\mathbf{v}, \mathbf{w}) - \mu_x(\mathbf{v})\mu_y(\mathbf{w}) \\&= \mathbb{E}_{\mathbf{xy}}[k(\mathbf{x}, \mathbf{v})l(\mathbf{y}, \mathbf{w})] - \mathbb{E}_{\mathbf{x}}[k(\mathbf{x}, \mathbf{v})]\mathbb{E}_{\mathbf{y}}[l(\mathbf{y}, \mathbf{w})], \\&= \text{cov}_{\mathbf{xy}}[k(\mathbf{x}, \mathbf{v}), l(\mathbf{y}, \mathbf{w})].\end{aligned}$$

- 1 Transforming $\mathbf{x} \mapsto k(\mathbf{x}, \mathbf{v})$ and $\mathbf{y} \mapsto l(\mathbf{y}, \mathbf{w})$ (from \mathbb{R}^{d_y} to \mathbb{R}).
- 2 Then, take the covariance.

The kernel transformations turn the linear covariance into a dependence measure.

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Alternative Form of $\hat{u}(\mathbf{v}, \mathbf{w})$

- Recall $\widehat{\text{FSIC}}^2 = \frac{1}{J} \sum_{i=1}^J \hat{u}(\mathbf{v}_i, \mathbf{w}_i)^2$
- Let $\widehat{\mu_x \mu_y}(\mathbf{v}, \mathbf{w})$ be an unbiased estimator of $\mu_x(\mathbf{v})\mu_y(\mathbf{w})$.
- $\widehat{\mu_x \mu_y}(\mathbf{v}, \mathbf{w}) := \frac{1}{n(n-1)} \sum_{i=1}^n \sum_{j \neq i} k(\mathbf{x}_i, \mathbf{v}) l(\mathbf{y}_j, \mathbf{w})$.
- An unbiased estimator of $u(\mathbf{v}, \mathbf{w})$ is

$$\begin{aligned}\hat{u}(\mathbf{v}, \mathbf{w}) &= \hat{\mu}_{xy}(\mathbf{v}, \mathbf{w}) - \widehat{\mu_x \mu_y}(\mathbf{v}, \mathbf{w}) \\ &= \frac{2}{n(n-1)} \sum_{i < j} h_{(\mathbf{v}, \mathbf{w})}((\mathbf{x}_i, \mathbf{y}_i), (\mathbf{x}_j, \mathbf{y}_j)),\end{aligned}$$

where

$$h_{(\mathbf{v}, \mathbf{w})}((\mathbf{x}, \mathbf{y}), (\mathbf{x}', \mathbf{y}')) := \frac{1}{2} (k(\mathbf{x}, \mathbf{v}) - k(\mathbf{x}', \mathbf{v})) (l(\mathbf{y}, \mathbf{w}) - l(\mathbf{y}', \mathbf{w})).$$

- $\hat{u}(\mathbf{v}, \mathbf{w})$ is a one-sample 2^{nd} -order U-statistic, given (\mathbf{v}, \mathbf{w}) .

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Independence Test with HSIC [Gretton et al., 2005]

- Hilbert-Schmidt Independence Criterion.

$$\text{HSIC}(X, Y) = \text{MMD}(P_{xy}, P_x P_y) = \|u\|_{\text{RKHS}}$$

(need two kernels: k for X , and l for Y).

- Empirical witness:

$$\hat{u}(\mathbf{v}, \mathbf{w}) = \hat{\mu}_{xy}(\mathbf{v}, \mathbf{w}) - \hat{\mu}_x(\mathbf{v})\hat{\mu}_y(\mathbf{w})$$

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- $\text{HSIC}(X, Y) = 0$ if and only if X and Y are independent.
- Test statistic = $\|\hat{u}\|_{\text{RKHS}}$ (“flatness” of \hat{u}). Complexity: $\mathcal{O}(n^2)$.

Key: Can we measure the flatness by other way that costs only $\mathcal{O}(n)$?

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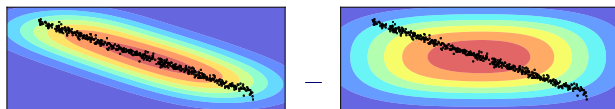
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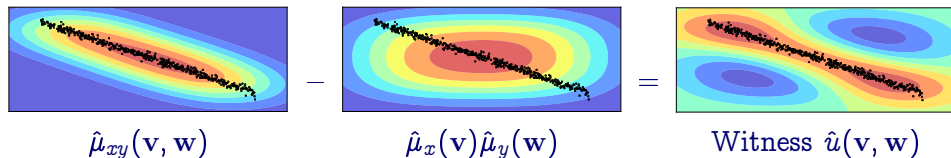
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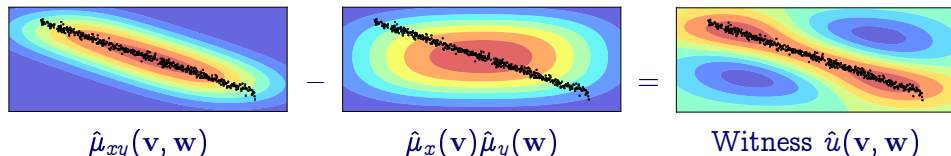
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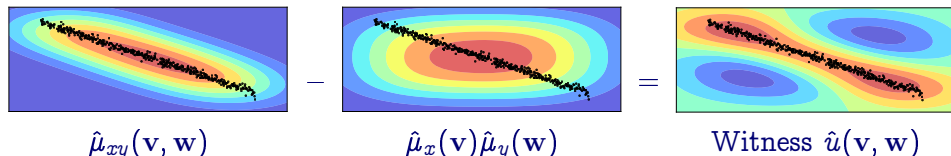
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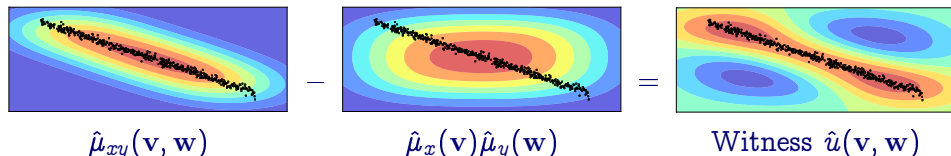
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Proposal: The Finite Set Independence Criterion (FSIC)

Idea: Evaluate $\hat{u}^2(\mathbf{v}, \mathbf{w})$ at only finitely many test locations.

■ A set of random J locations: $\{(\mathbf{v}_1, \mathbf{w}_1), \dots, (\mathbf{v}_J, \mathbf{w}_J)\}$

■ $\widehat{\text{FSIC}}^2(X, Y) = \frac{1}{J} \sum_{i=1}^J \hat{u}^2(\mathbf{v}_i, \mathbf{w}_i)$

■ Complexity: $\mathcal{O}((d_x + d_y)Jn)$. Linear time.

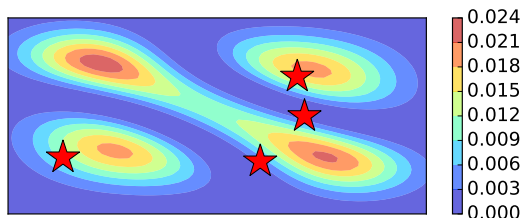
■ Can $\widehat{\text{FSIC}}^2(X, Y) = 0$ even if X and Y are dependent??

Proposal: The Finite Set Independence Criterion (FSIC)

Idea: Evaluate $\hat{u}^2(\mathbf{v}, \mathbf{w})$ at only finitely many test locations.

- A set of random J locations: $\{(\mathbf{v}_1, \mathbf{w}_1), \dots, (\mathbf{v}_J, \mathbf{w}_J)\}$

- $\widehat{\text{FSIC}}^2(X, Y) = \frac{1}{J} \sum_{i=1}^J \hat{u}^2(\mathbf{v}_i, \mathbf{w}_i)$



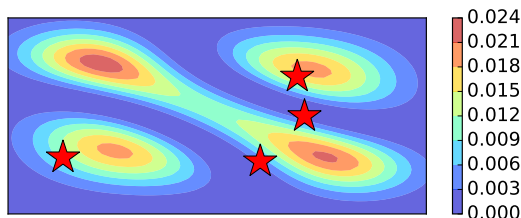
- Complexity: $\mathcal{O}((d_x + d_y)Jn)$. Linear time.
- Can $\widehat{\text{FSIC}}^2(X, Y) = 0$ even if X and Y are dependent??

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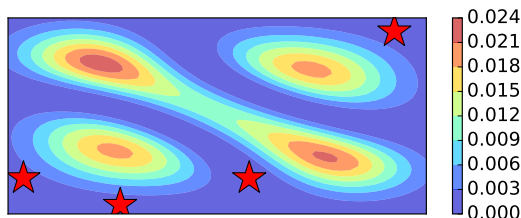
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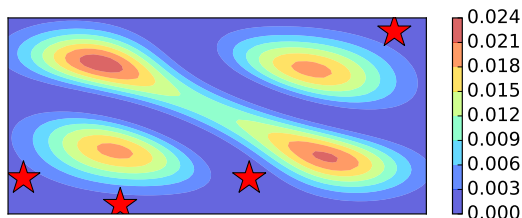
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- Can $\text{FSIC}^2(X, Y) = 0$ even if X and Y are dependent??
- No. Population $\text{FSIC}(X, Y) = 0$ iff $X \perp Y$, almost surely.

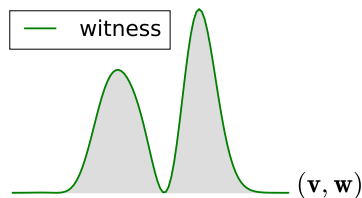
HSIC vs. FSIC

Recall the witness

$$\hat{u}(\mathbf{v}, \mathbf{w}) = \hat{\mu}_{xy}(\mathbf{v}, \mathbf{w}) - \hat{\mu}_x(\mathbf{v})\hat{\mu}_y(\mathbf{w}).$$

HSIC [Gretton et al., 2005]

$$= \|\hat{u}\|_{\text{RKHS}}$$

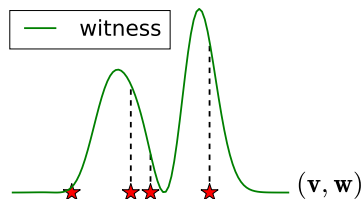


Good when difference between p_{xy} and $p_x p_y$ is spatially diffuse.

- \hat{u} is almost flat.

FSIC [proposed]

$$= \frac{1}{J} \sum_{i=1}^J \hat{u}^2(\mathbf{v}_i, \mathbf{w}_i)$$



Good when difference between p_{xy} and $p_x p_y$ is local.

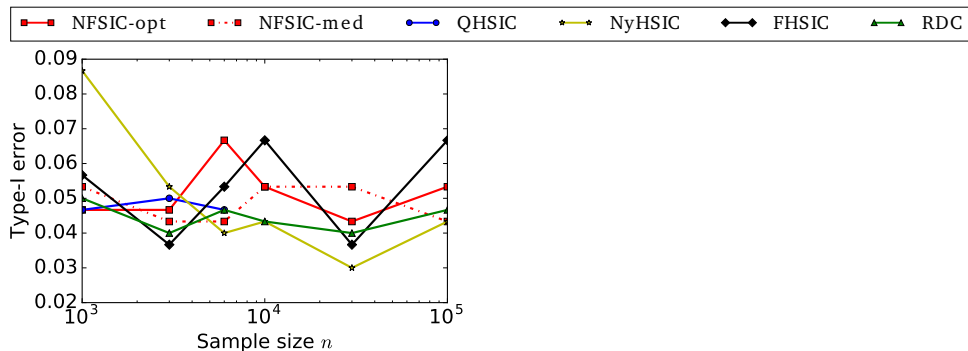
- \hat{u} is mostly zero, has many peaks (feature interaction).

Toy Problem 1: Independent Gaussians

- $X \sim \mathcal{N}(\mathbf{0}, \mathbf{I}_{d_x})$ and $Y \sim \mathcal{N}(\mathbf{0}, \mathbf{I}_{d_y})$.
- Independent X, Y . So, H_0 holds.
- Set $\alpha := 0.05$, $d_x = d_y = 250$.

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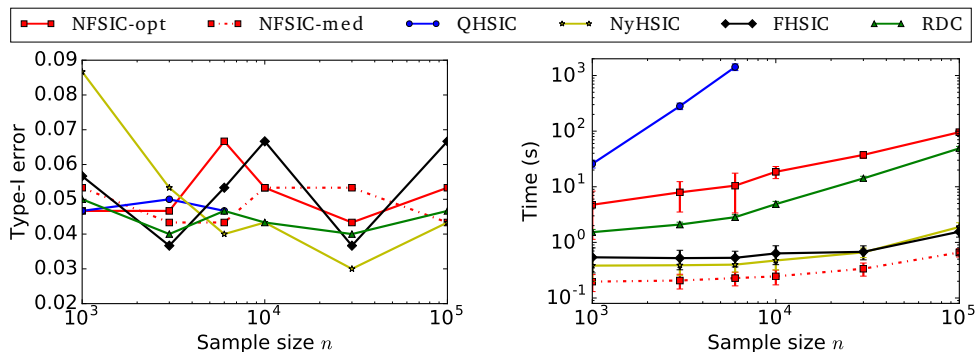
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- Correct type-I errors (false positive rate).

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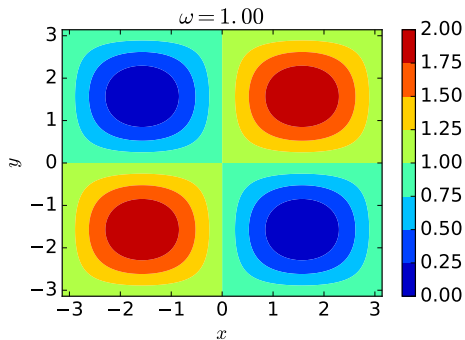
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Toy Problem 2: Sinusoid

- $p_{xy}(x, y) \propto 1 + \sin(\omega x) \sin(\omega y)$ where $x, y \in (-\pi, \pi)$.
- Local changes between p_{xy} and $p_x p_y$.
- Set $n = 4000$.

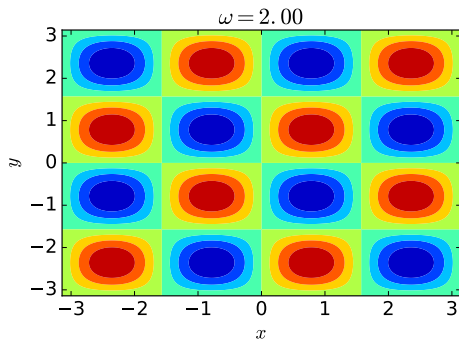
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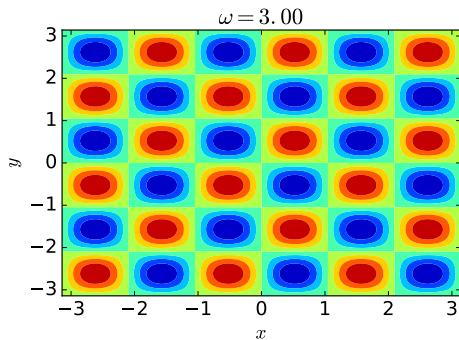
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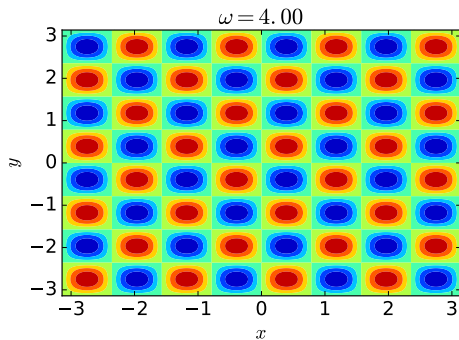
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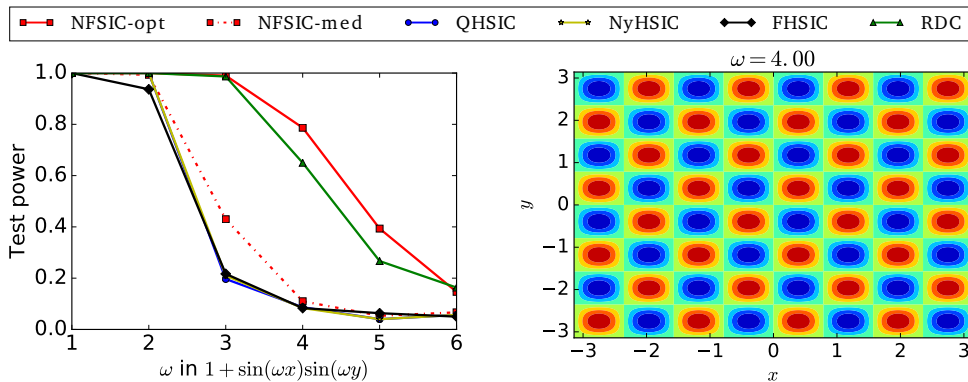
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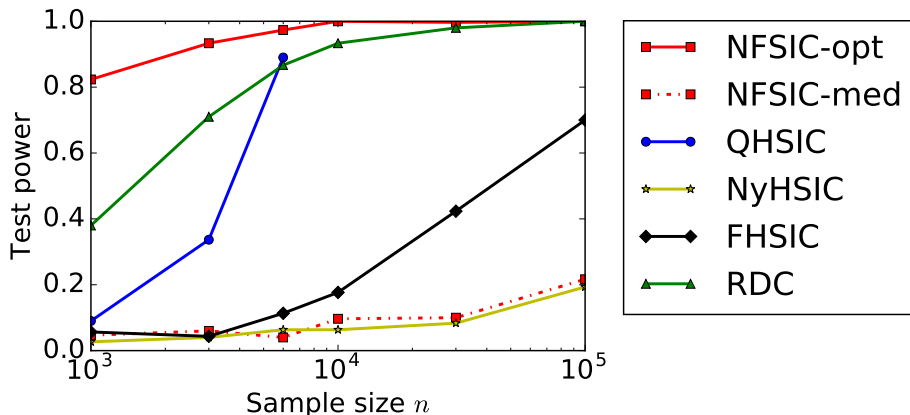
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Main Point: NFSIC can handle well the local changes in the joint space.

Toy Problem 3: Gaussian Sign

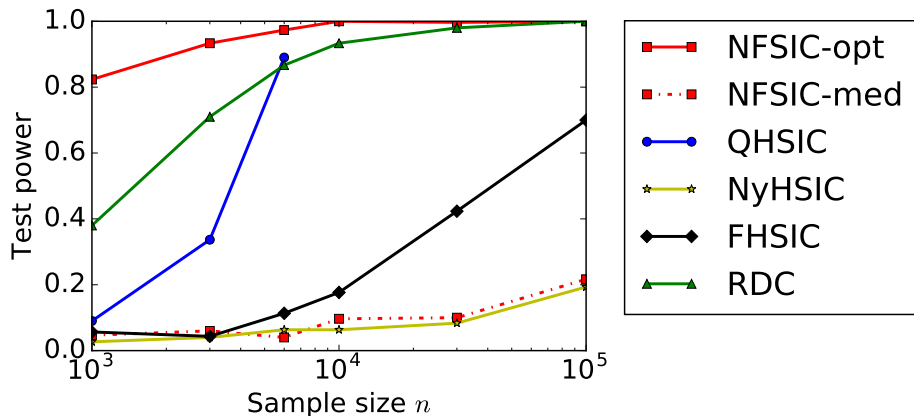
- $y = |Z| \prod_{i=1}^{d_x} \text{sign}(x_i)$, where $\mathbf{x} \sim \mathcal{N}(\mathbf{0}, \mathbf{I}_{d_x})$ and $Z \sim \mathcal{N}(0, 1)$ (noise).
- Full interaction among x_1, \dots, x_{d_x} .
- Need to consider all x_1, \dots, x_d to detect the dependency.



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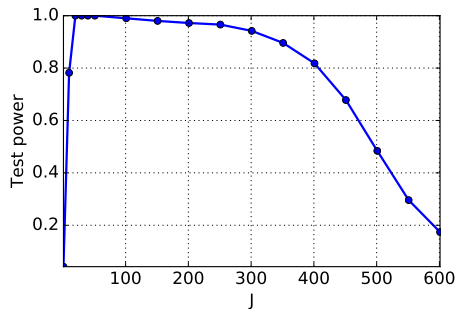
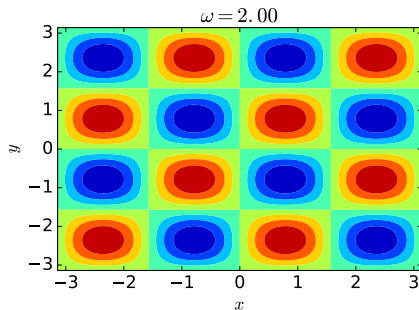
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Test Power vs. J

- Test power *does not* always increase with J (number of test locations).
- $n = 800$.



- Accurate estimation of $\hat{\Sigma} \in \mathbb{R}^{J \times J}$ in $\hat{\lambda}_n = n\hat{\mathbf{u}}^\top (\hat{\Sigma} + \gamma_n \mathbf{I})^{-1} \hat{\mathbf{u}}$ becomes more difficult.
- Large J defeats the purpose of a linear-time test.

Real Problem: Million Song Data

Song (X) vs. year of release (Y).

- Western commercial tracks from 1922 to 2011 [Bertin-Mahieux et al., 2011].
- $X \in \mathbb{R}^{90}$ contains audio features.
- $Y \in \mathbb{R}$ is the year of release.

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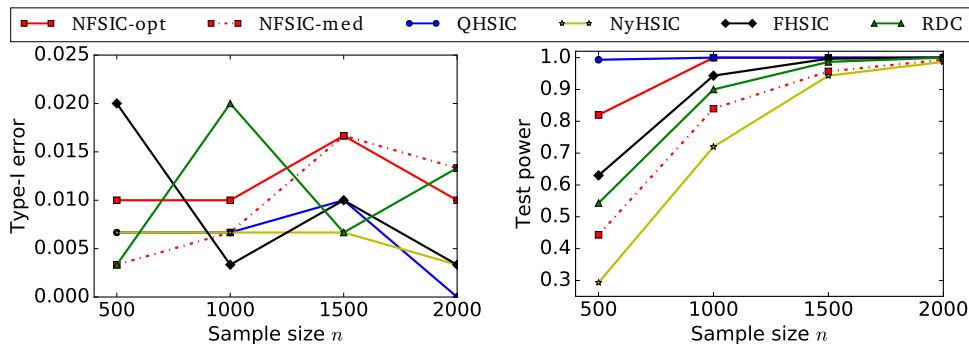
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- Break (X, Y) pairs to simulate H_0 .

NFSIC-opt has the highest power among the linear-time tests.

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